ABSTRACT

Objective  To measure the effects of faecal immunochemical test (FIT) for colorectal cancer (CRC) screening on overall and site-specific long-term effectiveness of population-based organised service screening.

Design  A prospective cohort study of Taiwanese nationwide biennial FIT screening was performed. A total of 5,417,699 eligible subjects were invited to attend screening from 2004 through 2009 and were followed up until 2014. We estimated the adjusted relative rates (aRRs) on the effectiveness of reducing advanced-stage CRC (stage II+) and CRC death by Bayesian Poisson regression models with the full adjustment for a cascade of self-selection factors (including the screening rate and the colonoscopy rate) and the completeness of colonoscopy together with demographic features.

Results  FIT screening (exposed vs unexposed) reduced the incidence of advanced-stage CRC (48.4 vs 75.7 per 100,000) and mortality (20.3 vs 41.3 per 100,000). Statistically significant reductions of both incidence of advanced-stage CRCs (aRR = 0.66, 95% CI 0.63 to 0.70) and deaths from CRC (aRR = 0.60, 95% CI 0.57 to 0.64) were noted. FIT screening was more effective in reducing distal advanced-stage CRCs (aRR = 0.61, 95% CI 0.58 to 0.64) and CRC mortality (aRR = 0.56, 95% CI 0.53 to 0.60) than proximal advanced CRCs (aRR = 0.84, 95% CI 0.77 to 0.92) and CRC mortality (aRR = 0.72, 95% CI 0.66 to 0.80).

Conclusion  A large-scale population-based biennial FIT screening demonstrates 34% significant reduction of advanced-stage CRCs and 40% reduction of death from CRC with larger long-term effectiveness in the distal colon than the proximal colon. Our findings provide a strong and consistent evidence-based policy for supporting a sustainable population-based FIT organised service screening worldwide. The disparity of site-specific long-term effectiveness also provides an insight into the remedy for lower effectiveness of FIT screening in the proximal colon.

INTRODUCTION

Colorectal cancer (CRC) is currently the fourth leading cause of cancer death worldwide, and screening is effective in reducing its incidence and mortality.1 While endoscopy-based screening is the most effective method, stool-based screening tests specifically identify high-risk subjects for colonoscopy, rendering it the most popular screening test worldwide, especially in regions where the clinical capacity of endoscopy is limited. Faecal immunochemical test (FIT) has been reported to outperform guaiac faecal occult blood tests.2 The effectiveness of FIT in reducing CRC mortality...
has been demonstrated recently by several population-based studies.3–6 Despite these observational studies, there is an urgent need for health decision-makers worldwide to evaluate the long-term effectiveness of FIT screening in each large-scale organised service screening in order to ensure that its quality assurance is as good as that of these previous organised CRC screenings that have completely or partially used the FIT test and have already demonstrated the significant effectiveness of mortality reduction.4–6 So doing also provides an evidence-based policy for FIT screening with sustainability in order to meet the Sustainable Development Goal (SDG) 3 of global health.7 Our previous report on the results from the inaugural 5 years of the Taiwan Colorectal Cancer Screening Program with only one-fifth of screening rate has already demonstrated a significant 10% reduction in CRC mortality, almost attaining the performance of guaiac test with a longer follow-up time of more than 10 years.3,4 While our nationwide programme has been expanded to around 60% screening rate, it provides an opportunity to demonstrate whether the long-term effectiveness of population-based organised FIT service screening can be achieved like colonoscopy screening evaluated with a randomised trial design or a modelling approach.5–10

Moreover, as previous studies have demonstrated that FIT is less sensitive for detecting proximal advanced neoplasm than distal ones,11 12 whether the long-term effectiveness of FIT in reducing mortality from CRC and advanced-stage CRC also varies with anatomical site is therefore worthy of being investigated. To estimate the unbiased overall and site-specific effectiveness in population-based FIT service screening as if obtained in a randomised controlled trial with intention-to-treat (ITT) analysis, one has to make allowance for relevant factors pertaining to self-selection bias and incomplete colonoscopy.

In the present study, we aimed to elucidate whether and how the Taiwanese nationwide population-based FIT screening can lead to the reduction of overall and site-specific mortality and incident advanced-stage CRC by analysing a large-scale cohort eligible for screening in the inaugural 5 years of the programme with continuous biennial FIT screening with 10 years of follow-up. Such an evaluation would be based on a newly developed method for the full adjustment for relevant characteristics that affect self-selection bias and the completeness of colonoscopy.

METHODS
Taiwan Colorectal Cancer Screening Program
Taiwan Colorectal Cancer Screening Program is a nationwide population-based FIT service screening launched in 2004. The details of this screening programme have been described in full elsewhere.3 In brief, a biennial single-sample FIT screening has been provided universally for subjects aged 50–69 years since the inaugural period (2004–2009). In addition to age range, criteria also included whether the enrolled subject is free of alarming symptoms of CRC such as bowel habit change, defecation with mucous or passage of blood, and tenesmus, all of which were assessed by primary healthcare providers (physician, nurse and public health personnel). It is mandatory for them to obtain and report the signed record on the result of the eligibility assessment in written form. Information on the screened cohort has been centralised in our mass-screening registry. As mass-screening registry and population registry have been stored as the centralised database in a nationwide scale, non-participants were ascertained after the linkage of the screened cohort from mass-screening registry with all eligible subjects recorded in population registry. During this period, people could only obtain the FIT kit at regional health centres prompted by the designated awareness campaign. To boost up the screening coverage, the government started to roll out the programme in 2010. People could also have the uptake of FIT screening in hospitals or clinics, in addition to the aforementioned sites, and the way of assessing eligibility was identical.13 One of two separate FIT kits (OC-SENSOR, Eiken Chemical Ltd, Tokyo, Japan, or HM-JACK, Kyowa Medex Co, Tokyo, Japan) was selected by each municipality or hospital/clinic according to its own purchasing process, both using 20 µg of haemoglobin/g of faeces as the cut-off to determine positivity. People were notified of the FIT results by postal mail or by physician in the outpatient clinics, and those who had positive FITs were then referred for colonoscopy as a diagnostic examination within 3 months, and those who were considered as infeasible for colonoscopy or failed complete colonoscopy were then offered double-contrast barium enema as an alternative diagnostic procedure. This process was intensified by sending second or third notices via public health workers or nursing staff networks in the regional health centres, hospitals or clinics. A cascade of the entire screening process from uptake screening until diagnostic examination with complete colonoscopy to reach the cecum is diagrammed in figure 1. All these relevant screening characteristics (such as the screening rate, the FIT positivity rate, the referral rate of diagnostic examinations, the rate of selecting colonoscopy as diagnostic examination, complete colonoscopy rate, and detection rates for advanced adenoma and cancers) of individual municipalities were regularly monitored using the central screening database in the Health Promotion Administration of the Taiwanese government. It should be noted that although these characteristics have improved with time (complete colonoscopy rate increased from 75% in 2004 until 96% in 2014), they have to be adjusted during the inaugural 5 years in order to estimate the unbiased effectiveness of FIT screening as close as to that estimated with a randomised controlled design while the self-selection bias and the completeness of colonoscopy are taken into account (see further). Those who had neoplasms detected and resected at colonoscopy were recommended for subsequent surveillance colonoscopy based on the initial findings as recommended by the current surveillance guidelines, and those who had negative examination were advised to receive biennial FIT screening within the programme. The FIT laboratories and endoscopic units are periodically audited and accredited by the screening organiser, which is described in our previous publications.14 15

Study design, population and data collection
In this prospective cohort study, 5 417 699 subjects who were considered as eligible for CRC screening during the inaugural 5 years (2004–2009) were continuously offered the subsequent screen on biennial basis and followed up until the end of 2014 (figure 2).3 Participants who underwent at least one FIT screening during the period of 2004–2014 constituted the exposed group, and the rest of the population formed the unexposed group. Note that whether the exposed group was further referred to have diagnostic examination and whether to have complete colonoscopy would be refined according to figure 1, when allowance was made for self-selection bias and the quality of complete colonoscopy. Events were defined as incident advanced-stage (stage II+) CRC and CRC deaths. The screening database was linked to the national cancer registry and national death registry, from which the two aforementioned events could
be ascertained and verified. The coverage rate of the national cancer registry was reported to be 98.6% with accuracy of greater than 99%. Both incident CRCs and CRC deaths were ascertained with the follow-up until the end of 2014. CRCs were staged in the light of the sixth or seventh version of the American Joint Committee on Cancer staging system. Colon anatomical site above the level of splenic flexure is defined as the proximal colon and the remaining segments and rectum as the distal colon.

**Statistical analysis**

Descriptive results for FIT positivity rate, diagnostic examination rate and detection rate for CRC are reported as percentages. The anatomical site and stage distribution of CRCs were identified in the exposed and the unexposed groups during the study period. FIT interval cancers were CRCs occurring within 2 years after negative FIT were defined, which enables us to calculate the programme sensitivity by anatomical site. Colonoscopy interval cancers refer to the incident CRCs that become symptomatic and diagnosed within the recommended surveillance interval after the baseline colonoscopy without the diagnosis of CRC. Incidence and mortality rates were expressed as events per 10^5 person-years. Because it is a population-based service screening, subjects were enrolled by a gradual rolling-out strategy that constitutes a prospective cohort with the staggered entry to screen in different calendar years for all eligible Taiwanese population aged 50–69 years as shown in figure 2. This also means that the screening rate increased with year, although the overall
screening rate was up to 57% until 2014. The calculation of person-years for each eligible individual was therefore based on the follow-up time from the date of the starting year (2004) until the end of follow-up, which was defined as the occurrence of an event, death from competing causes, age of 69 years or the end of the study period, whichever came earlier. To classify person-years into two groups, the exposed group (attender) and the unexposed group (non-attender), individual variation with time to attend the first screen (time-dependent property) should be considered, namely, those who had been unexposed to screen since 2004 but attended FIT screening later during the study period contributed their person-years to the unexposed group prior to attend FIT screening, but person-years thereafter were attributed to the exposed group. Such a dynamic change of person-years belonging to the two groups due to the eligible subjects attending the first screen in a staggered entry is considered in the following Bayesian Poisson regression model for correcting self-selection bias with adjustment for the screening rate in a time-dependent manner. The same calculation of person-years accompanied with the exposure status to attend the screen was also applied to more than two groups further classified by referral versus non-referral to diagnostic examination, colonoscopy versus non-colonoscopy diagnostic examination, and complete versus incomplete colonoscopy while the following full adjustment for correcting self-selection bias and the completeness of colonoscopy was made.

**Self-selection bias adjustment with the screening rate**

As the screening rate of population-based service screening with a staggered entry as mentioned previously is a reflection of aggregated self-selection on population level, it is necessary to estimate unbiased effectiveness between the exposed group and the unexposed group like the comparison between the invited and the uninvited group in the absence of screening using the language of a randomised controlled trial design with ITT analysis. This is exemplified in online supplemental materials (see the online supplemental equation (S-1)). The unbiased relative risk for measuring the reduction of CRC death as a result of FIT screening is expressed as

\[
\frac{P(CRC\ Death\ (0)\mid\ Invited\ (1))}{P(CRC\ Death\ (0)\mid\ Uninvited\ (0))}
\]

To approximate this unbiased estimate on the population level, we have to correct self-selection bias with adjustment for the screening rate extending the method that has been developed and widely used in the evaluation of the effectiveness of a population-based organised service screening programme. The risk of being dead from CRC as in the numerator of the equation (1) can be decomposed into the exposed group and the unexposed group pursuant to ITT analysis. Note that relative risk is changed to relative rate when person-year is used for the denominator. Take the primary endpoint of CRC mortality for example, the first adjusted relative rate (aRR) in relation to the screening rate can be derived with the following expression:

\[
aRR_e = (\text{screening rate}) \times \text{RR}_e + (1 - \text{screening rate}) \times \text{RR}_\tau
\]

where \( \text{RR}_e \) is the mortality rate of CRC in the exposed group compared with the mortality rate in the uninvited group, and \( \text{RR}_\tau \) is the mortality rate of CRC in the unexposed group compared with the mortality rate in the uninvited group. The two relative comparisons, the exposed vs the uninvited group (\( \text{RR}_e \)) and the unexposed group vs the uninvited group (\( \text{RR}_\tau \)), are weighted by the screening rate. Note that the uninvited group (theoretically used in a randomised control design) can be approximated by using the pre-screening period with adjustment for an increasing growth rate (see below). The details of derivation are given in online supplemental materials.

**Correction of self-selection bias with full adjustment**

As self-selection bias can also be affected by the referral rate of diagnostic examinations among FIT positives and the choice of colonoscopy (colonoscopy rate) as diagnostic examination, and the quality of colonoscopy regarding the completeness of colonoscopy (cecal intubation rate), particularly in the inaugural period, the equation (2) following the equation (1) is further expanded to get the second adjusted relative rate (aRR) with the full adjustment for a cascade of these factors with the following expression:

\[
aRR_e = \frac{\text{Referral rate} \times \text{Cecal intubation rate} \times \text{Colonoscopy rate} \times \text{Screening rate}}{\text{Referral rate} \times \text{Screening rate}} + (1 - \text{screening rate}) \times \text{RR}_\tau
\]

\[
(\text{Part I}) = (\text{Cecal intubation rate}) \times (\text{Colonoscopy rate}) \times (\text{Screening rate})
\]

\[
\text{Referral rate} \times \text{Screening rate}
\]

\[
(\text{Part II}) = (1 - \text{Cecal intubation rate}) \times (\text{Colonoscopy rate}) \times (\text{Screening rate})
\]

\[
\text{Referral rate} \times \text{Screening rate}
\]

\[
(\text{Part III}) = (1 - \text{Colonoscopy rate}) \times (\text{Referral rate}) \times \text{Screening rate}
\]

\[
(\text{Part IV}) = (1 - \text{Referral rate}) \times (\text{Positive rate}) \times \text{Screening rate}
\]

\[
(\text{Part V}) = (1 - \text{Positive rate}) \times \text{Screening rate}
\]

\[
\text{RR}_{e}, \text{RR}_{\tau}, \text{RR}_{\text{CT}}, \text{RR}_{\text{CF}} \text{ and RR}_{\text{C}} \text{ represent relative rates for a series of relative comparisons, including those with complete colonoscopy (successful cecal intubation), incomplete colonoscopy, not selecting colonoscopy as diagnostic examination, refusing diagnostic examination and negative FIT results compared with the uninvited group as indicated in equation (1). All these relative rates are weighted by a cascade of screening characteristics affecting the performance of the screening process denoted by \( \text{RR}_{\text{e}} \) for the exposed group, \( \text{RR}_{\text{OF}} \) for positive FIT test, \( \text{RR}_{\text{UF}} \) for the referral rate for diagnostic examination, \( \text{RR}_{\text{CF}} \) for the colonoscopic rate and \( \text{RR}_{\text{C}} \) for the rate of completing colonoscopy with a reach to caecum. The derivation of equation (3) is elaborated in the online supplemental materials.}

**Bayesian Poisson regression model**

To make the two adjusted relative rates (aRRs) amenable to estimation, we exploited the Bayesian Poisson regression model with time-varying person-years based on equation (2) and also online supplemental equation (S-3) in the online supplemental materials, with adjustment for the screening rate alone written as follows:

\[
\log(\text{expected number of advanced} - \text{stage CRC or death from CRC}) = \log(\text{time} - \text{dependent person} - \text{years}) + \beta_1 + \beta_2 \cdot \text{exposed} + \beta_3 \cdot \text{age} + \beta_4 \cdot \text{gender}
\]

where \( \beta_1 (0.0443=\log(1.045)) \) is the natural growth rate of incidence trend of CRC as indicated earlier. The Bayesian Poisson regression model estimated this aRR through two relative rates (\( \text{RR}_{\text{e}} \) and \( \text{RR}_{\text{C}} \)) as indicated previously in terms of the exponent of two regression coefficients, \( \beta_1 \) and \( \beta_2 \), obtained by modelling the relationship of the status of screen (1=exposed for attenders, 0=unexposed for non-attenders) represented by two indicator functions to the number of CRC death.

Equation (4) takes into account extraneous factors, age, gender and increasing incidence rate, that have been applied previously. In Taiwan, as incidence of CRC is increasing annually, 4.5% of the growth rate of biologically increasing incidence rate is therefore
required to make adjustment by extrapolation based on data from 1998 to 2003 obtained from the national cancer registry before the launch of screening. Thus, the annual mortality of 50.2 per 100 000 between 1998 and 2003 was adjusted to 52.5 per 100 000, which is taken as the expected mortality in the absence of mass screening that is supposed to be equivalent to the uninvited (control) group of a randomised controlled trial for selection bias adjustment.

Regarding the full adjustment for self-selection bias and the completeness of colonoscopy, the corresponding Bayesian Poisson regression model expanded from equation (3) is expressed as

\[ \log(\text{expected number of advanced stage CRC or death from CRC}) = \log(\text{time - dependent person - years}) + \beta_0 + \beta_1 \cdot (\text{Complete colonoscopy}) + \beta_2 \cdot (\text{Incomplete colonoscopy}) + \beta_3 \cdot (\text{Non - colonoscopy exam}) + \beta_4 \cdot (\text{Non - referral}) + \beta_5 \cdot (\text{FIT - negative}) + \beta_6 \cdot (\text{unexposed}) + \beta_7 \cdot (\text{age}) + \beta_8 \cdot (\text{gender}) \]  

(5)

Recall that the distinction of the Bayesian Poisson regression model shown in equation (5) different from that in the equation (4) is that the equation (5) further divides the exposed group in the equation (4) according to a cascade of factors after the exposure to screen (including positive FIT, referral rate, the choice of colonoscopy and complete colonoscopy), as shown in figure 1 with the corresponding coefficients from \( \beta_1 \) to \( \beta_8 \). The details of the derivation are also elaborated in the online supplemental materials.

The mortality of CRC and the incidence of advanced-stage CRC reduction after adjustment was calculated as (1 - aRR1 or 2) \times 100%.

Patient and public involvement
It was not appropriate or possible to involve patients or the public in the design, conduction, reporting or dissemination plans of our research.

RESULTS
Of 5 417 699 subjects aged 50–69 years enrolled in our prospective cohort as mentioned previously, a total of 3 067 853 subjects did not receive any FIT screening, yielding a screening coverage rate of 52.3%. The positivity rate of FIT was 7.0% in the first round or more rounds of FIT screening, indicating a repeat screening rate of 52.3%. The positivity rate of FIT was 7.0% in the first round and 62.5% in the subsequent rounds (94.1% and 6.4% in the subsequent rounds). A total of 70.0% of FIT positive cohort as mentioned previously, a total of 3 067 853 subjects

<table>
<thead>
<tr>
<th>Table 1 Characteristics of the study cohort</th>
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<td>CRC death, n</td>
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<td>CRC, colorectal cancer.</td>
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Cancer stage distribution between screened and unscreened groups
The proportions of CRC identified in the exposed group and the unexposed group at each stage within the study period are shown in table 2. The rate of incomplete cancer stage information in detected cancers was 10.8% in the exposed group and 15.9% in the unexposed group.

Incidence of advanced-stage CRC after FIT screening
Table 3 shows the incidence rate of advanced-stage CRC, which was 48.4 and 75.7 per 10⁵ person-years in the exposed and the unexposed groups, respectively. The crude relative rate (cRR) of incidence of advanced-stage CRC was 0.64 (95% CI 0.62 to 0.66). Table 3 also shows after adjustment for age, gender and the screening rate in relation to self-selection bias, using equation (2) and the corresponding regression equation (4), the aRR of reducing incident advanced-stage CRCs was 0.71 (95% CI 0.68 to 0.75). Further adjustment for referral and completeness of colonoscopy using equations (3) and (5) gave 0.66 (95% CI 0.63 to 0.70) of the aRR of reducing incident advanced-stage CRCs.

CRC mortality after FIT screening
Table 3 shows CRC mortality was 20.3 and 41.3 per 10⁵ person-years in the exposed and the unexposed groups, respectively. The cRR of CRC deaths was 0.49 (95% CI 0.47 to 0.51). After adjusting for age, gender and the screening rate related to self-selection bias, the aRR of reducing death from CRC was 0.65 (95% CI 0.62 to 0.69). The full adjustment for selection bias and completeness of colonoscopy involving the performance of a cascade of screening process led to an estimate of 0.60 (95% CI 0.57 to 0.64) of the aRR of reducing death from CRC.

Incidence of advanced-stage CRC and mortality by anatomical sites
Table 3 also shows the aRR values for the reduction of site-specific incidence of advanced-stage CRCs with only adjustment for the screening rate were 0.92 (95% CI 0.84 to 1.00) and 0.65 (95% CI 0.62 to 0.69) for the proximal and the distal colon, respectively. After the full adjustment, the aRR values of reducing incident advanced-stage CRC were 0.84 (95% CI 0.77 to 0.92) and 0.61 (95% CI 0.58 to 0.64) for the proximal and the distal colon, respectively. The corresponding aRR values for the reduction of CRC mortality with only adjustment for the screening rate were 0.79 (95% CI 0.72 to 0.87) and 0.61 (95% CI 0.57 to 0.64), respectively. The counterparts of reducing CRC mortality with the full adjustment were 0.72 (95% CI 0.66 to 0.80) and 0.56 (95% CI 0.53 to 0.59), respectively. It is obvious that the incremental effectiveness was larger in the proximal colon (8% for incidence of advanced-stage CRC and 7% for CRC mortality) than in the distal colon (4% for incidence of advanced-stage CRC and 5% for CRC mortality) after full adjustment from the screening rate to the completeness of colonoscopy.

The sensitivity analysis also shows that the aRR of reducing death from CRC mortality provided that 100% completeness of colonoscopy can be achieved.

Site-specific accuracy of FIT
The disparity of site-specific finding on effectiveness was also supported by the difference in site-specific sensitivity with the order
of 62.9% for proximal cancers (72.7% for advanced stage ones) and 73% for distal cancers (83.2% for advanced stage ones).

DISCUSSION
This is the first study with the evaluation of long-term effectiveness of population-based FIT service screening based on the full adjustment for a cascade of self-selection factors from screening uptake, the confirmatory exam rate, to the choice of colonoscopy and also the allowance made for the completeness of colonoscopy. This supports the site-specific reduction in incidence of CRC in the proximal colon compared with the distal colon particularly when the allowance was made for the completeness of colonoscopy.

As far as the former is concerned, our site-specific findings found that larger effectiveness was gained for the proximal colon than the distal ones. This is supported by the fact that stage distribution was very different for proximal and distal CRC, which might in part account for the survival differences, and failed cecal intubation.\(^{14,21-24}\) It is therefore possible that the colonoscopy for detecting CRCs may vary with anatomical sites.

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<th>AJCC staging*</th>
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</table>

*Staging was based on AJCC sixth version before 2010 and seventh version thereafter. Stage information was incomplete in 10.8% in the screened group and 15.9% in unscreened group.
†Also includes FIT interval cancer, colonoscopy interval cancer and CRC of non-compliance with colonoscopy, in addition to CRC detected at first and subsequent screening rounds of next two columns.

AJCC, American Joint Committee on Cancer; CRC, colorectal cancer; FIT, faecal immunochemical test.
Second, CRCs even at the same stage with BRAF mutation (serrated pathway) were reported to have more grave outcome than CRC arisen via the so-called conventional pathway.23 26 Such a biological difference can also explain the difference in the completeness of colonoscopy.14 21–24 29 The disparity of site-specific sensitivity results on invasive CRCs also support FIT interval cancer, colonoscopy interval cancers, or CRC in those who were not compliant with diagnostic exams.14 21–24 29 The disparity of site-specific sensitivity estimates is consistent with the finding in a recent Italian study that proportional interval cancer rate was significantly higher in exposed group than in the unexposed group for calculating incidence of advanced cancer; person-years were 13 196 865 in the exposed group and 31 294 921 in the unexposed group for calculating CRC mortality.

| Table 3 | Incidence of advanced CRCs or mortality from CRC in the exposed and unexposed groups and their crude and adjusted relative rates of the effectiveness of FIT screening between the two groups |
|-----------------|-----------------|-----------------|-----------------|
|                | Unexposed group | Exposed group   |                    |
|                | (N=2 349 846)   | (N=3 067 853)   | Relative rate (95% CI), exposed versus unexposed |
|                | Case (n)        | Rate (per 1000) | Case (n)        | Rate (per 1000) | Crude | Age-adjusted and gender-adjusted | aRR1 | (full adjustment)2 |
| Incidence of advanced cancer | Overall 23 689 | 75.7           | 6381            | 48.4           | 0.64 (0.62 to 0.66) | 0.52 (0.50 to 0.53) | 0.71 (0.68 to 0.75) | 0.66 (0.63 to 0.70) |
|                        | Proximal 6127   | 19.6           | 2070            | 15.7           | 0.80 (0.76 to 0.84) | 0.62 (0.59 to 0.65) | 0.92 (0.84 to 1.00) | 0.84 (0.77 to 0.92) |
|                        | Distal 17 404   | 55.6           | 4267            | 32.3           | 0.58 (0.56 to 0.60) | 0.48 (0.46 to 0.50) | 0.65 (0.62 to 0.69) | 0.61 (0.58 to 0.64) |
| Mortality             | Overall 15 550  | 41.3           | 3077            | 20.3           | 0.49 (0.47 to 0.51) | 0.44 (0.43 to 0.46) | 0.65 (0.62 to 0.69) | 0.60 (0.57 to 0.64) |
|                        | Proximal 4004   | 10.6           | 988             | 6.5            | 0.61 (0.57 to 0.66) | 0.54 (0.50 to 0.58) | 0.79 (0.72 to 0.87) | 0.72 (0.66 to 0.80) |
|                        | Distal 11 440   | 30.4           | 2076            | 13.7           | 0.45 (0.43 to 0.47) | 0.41 (0.39 to 0.43) | 0.61 (0.57 to 0.64) | 0.56 (0.53 to 0.59) |

Person-years were 13 196 865 in the exposed group and 31 294 921 in the unexposed group for calculating incidence of advanced cancer; person-years was 15 179 449 in the exposed group and 37 658 371 in the unexposed group for calculating CRC mortality.

Remedies for reducing the disparity of site-specific effectiveness
In addition to providing a high-quality complete colonoscopy to reduce the disparity of site-specific interval cancer and long-term effectiveness of two-tier FIT screening, certain approaches can also be considered to reduce the disparity of site-specific FIT interval cancers resulting from the inaccuracy of FIT in the proximal colon. These include shortening interscreening intervals, increasing stool sample numbers and lowering the cutoff for determining FIT positivity. It should be also noted that these approaches may enhance the odds of detecting advanced

Comparison with previous studies on FIT screening programme
Based on the explanation on the completeness of colonoscopy responsible for the disparity of site-specific long-term effectiveness as indicated previously, it would not be surprised to find that after the correction for self-selection bias with the full adjustment, our results of reducing incidence of advanced-stage CRC and long-term effectiveness of mortality reduction resulting from a population-based service screening were consistent with those the results of two previous studies from the Northern Italy FIT screening programme4 and the KPNC organised CRC screening programme with various screening modalities (including outreach FIT) that are with high performance of the uptake of screening test and colonoscopy. Such a consistency implies that we would further expect an empirical finding on even larger reduction of CRC mortality by the continued Taiwanese nationwide FIT screening programme provided that these basic screening characteristics and the completeness of colonoscopy can be improved to a comparable level to the aforementioned FIT screening programmes. This again suggests high screening rate and the completeness of colonoscopy cannot be overemphasised in the implementation of a large population-based service screening.
adenoma and early CRC, but may increase the demand for colonoscopy and put further stress on the currently constrained colonoscopy capacity. Applying more sensitive tests such as the multitarget stool DNA test may be another viable approach to improve test sensitivity, but its high cost, uncertain or possibly low public acceptance, and unknown long-term effectiveness are obvious barriers for implementing it in a large-scale population-based screen at the present time. Quality assurance of FIT and appropriate collection and management of stool sample is also indispensable for reducing the risk of FIT interval cancer, thereby maximising the effectiveness of FIT screening.

Our previous study has demonstrated that different FIT kit might have different performance in terms of positive predictive value for CRC, leading to different risks of FIT interval cancers. Hence, the HM-JACK kits have been replaced with new generation HM-JACKarc in our programme. The strength of the current study is twofold. An innovative methodology was developed and applied to estimating the unbiased relative rate of two primary outcomes following FITT analysis with the full adjustment for a cascade of self-selection factors and the completeness of colonoscopy. Such a comprehensive adjustment that has never been done before not only renders the estimated effectiveness in population-based service screening as close as possible to the true value based on a randomised controlled trial design but also suggests that the quality of colonoscopy, more specifically, the completeness of colonoscopy, makes a contribution to the disparity of site-specific long-term effectiveness.

The present study is not without limitation. First, the diagnostic examination rate was unsatisfactory in Taiwanese programme. While the rate was about 80% in the inaugural 5 years of the programme as previously reported, it declined along with the rapid rolling out process since 2010 when FIT positive cases expanded by more than three times. The similar findings of unsatisfactory referral rate for colonoscopy were also noted in two recent US studies. Currently, the rate has gradually recovered to higher than 70% under the effort of the screening organiser, healthcare providers, and professional societies. Nevertheless, we have observed a significant reduction of CRC mortality and incidence of advanced-stage CRC, and once such a rate could be further improved, together with high screening rate and the colonoscopy quality, then we can expect an even larger magnitude of screening effectiveness. Second, we did not evaluate the overall incidence rate in our study as did in the KPNC study because, to a greater extent, the follow-up time in this study is still too short due to the staggered entry property of the current study and, to a lesser extent, still a lower screening rate and insufficiently high colonoscopy compliance rate with the early cohort (between 2004 and 2009) that is used for analysis. The reduction of overall incidence rate may require a longer follow-up period due to the long dwelling time from adenoma to invasive cancer. One would anticipate a reduction of the incidence of CRC from the screening programme owing to the removal of adenoma. Mandel et al has demonstrated a 17% (95% CI 6% to 27%) reduction with the biennial fecal occult blood test (FOBT) in the Minnesota trial. It could be underestimated with a relatively shorter follow-up time in the present study. Moreover, a longer follow-up time would also be required to deal with the lead-time issue in the early phase of service screening programme. Nonetheless, a comparably shorter dwelling time from early to advanced-stage cancer makes it possible to evaluate the reduction in advanced-stage cancer at an earlier timing as we did in the current study. The further follow-up of the cohort to elucidate the long-term effectiveness of reducing overall incidence is mandatory.

In conclusion, we demonstrated that FIT screening after a long-term follow-up is effective in reducing the risk of advanced-stage CRC and its mortality, with effectiveness consistently stronger for distal CRCs compared with proximal ones. Our current results on long-term effectiveness, together with the findings from previous studies, may provide a strong and consistent evidence-based policy for supporting a sustainable population-based FIT organised service screening worldwide, which is very meaningful for attaining the SDG 3 of global health.

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Supplementary Materials: Self-selection bias adjustment and Bayesian Poisson regression model in population-based cancer service screening

Population-based cancer service screening is often faced with self-selection bias. Namely, those who had participated in the screening are different from those who had not in terms of basic characteristics such as socio-economic status. To adjust such a kind of self-selection bias, it is imperative to follow intention-to-treat (ITT) analysis under the framework of a randomized controlled trial (RCT) design. We herein begin with the illustration of ITT analysis with a hypothetical example of population-based fecal immunochemical test (FIT) screening with a RCT design as shown in the following Supplementary Figure 1.

Supplementary Figure 1. The conceptual diagram showing study design for evaluating the effectiveness of a population-based FIT screening with a RCT design.
Relative risk of measuring effectiveness with ITT analysis in population-based FIT screening

In order to assess whether colorectal cancer (CRC) screening with FIT can reduce the incidence of advanced-stage CRC or CRC mortality, the eligible population are randomized into two groups, the invited group (receive screen with FIT) and the uninvited group (without FIT), denoted by I (the invited group) and \( \bar{I} \) (the uninvited group) in the Supplementary Figure 1, respectively. We here use CRC death as an example but the similar derivation can be also applied to the advanced-stage CRC. Even in the RCT design, the invited group is also classified into two groups, the exposed group (attender) and the unexposed group (non-attender) according to attendance rate. The relative risk of CRC mortality between group I and group \( \bar{I} \) is expressed as:

\[
\frac{P(\text{CRC death in the invited group } I)}{P(\text{CRC death in the uninvited group } \bar{I})} = \frac{P(D|I)}{P(D|\bar{I})} \quad (S-1)
\]

which is used to unbiasedly measure the reduction of death from CRC in light of ITT analysis to reduce self-selection bias in light of per protocol analysis if the corresponding relative risk between the exposed group and the uninvited group is used because those who are willing to attend the screen would not be representative of the control group derived from the underlying population. Note that relative risk is changed to relative rate when person-years is used for the denominator.
Relative rate of measuring effectiveness with ITT analysis in population-based FIT service screening

Unlike the RCT design, there is lacking of the uninvited group (the control group $\bar{I}$ in Supplementary Figure 1) in the evaluation of population-based FIT service screening as all the eligible population are invited for FIT screening. The comparator used for evaluation often relies on the unexposed group (see Supplementary Figure 1). It is more prone to self-selection bias when one would like to use the relative rate of being dead from CRC between the exposed group (E) and the unexposed group ($\bar{E}$) to measure the effectiveness of FIT screening with the following expression:

$$\frac{P(\text{CRC death in the exposed group } E)}{P(\text{CRC death in the unexposed group } \bar{E})}$$

To adjust for such a kind of self-selection bias, one has to return to use the ITT analysis to get unbiased estimated effectiveness in population-based FIT service screening as if obtained from a RCT design by making use of screening characteristics from uptake screening, referral rate for diagnostic examination, until the cecal intubation rate with complete colonoscopy.

We begin with self-selection in relation to the uptake of screen (screening arte) alone. The numerator regarding the risk for being dead from CRC in the invited group of (S-1) can be decomposed into two parts, the exposed group and the unexposed
group with the mathematical formula expressed as:

\[
RR = \frac{P(\text{CRC death in the invited group } I)}{P(\text{CRC death in the uninvited group } \bar{I})} = \frac{P(\text{CRC death|Invited})}{P(\text{CRC death|Uninvited})}
\]

\[
= \frac{P(D|I)}{P(D|\bar{I})} = \frac{P(E|I)P(D|E,I)+P(E|\bar{I})P(D|E,\bar{I})}{P(D|\bar{I})}
\]

where D, E, \(\bar{E}\), I and \(\bar{I}\) represents dead from CRC, exposed, not exposed, invited, and not invited to FIT screening, respectively.

Since we know the status of exposure to screen, the invitation (I) conveys little information, which is so-called conditional independence that simplifies \(P(D|E, I)\) and \(P(D|\bar{E}, I)\) into \(P(D|E)\) and \(P(D|\bar{E})\). Consider \(P(E|I)\) and \(P(\bar{E}|I)\) as the screening rate (\(r_E\)) and the complementary of screening rate (1-\(r_E\))

The equation above is reduced to

\[
\frac{P(D|E)P(E|I)+P(D|\bar{E})P(\bar{E}|I)}{P(D|\bar{I})} = \frac{P(D|E)}{P(D|\bar{I})} \times r_E + \frac{P(D|\bar{E})}{P(D|\bar{I})} \times (1 - r_E)
\]

(S-2)

Let

\[
RR_E = \frac{P(\text{CRC death in the exposed group } E)}{P(\text{CRC death in the uninvited group } \bar{I})} = \frac{P(D|E)}{P(D|\bar{I})}
\]

and

\[
RR_{\bar{E}} = \frac{P(\text{CRC death in the exposed group } \bar{E})}{P(\text{CRC death in the uninvited group } \bar{I})} = \frac{P(D|\bar{E})}{P(D|\bar{I})}
\]

be defined as two relative rates of CRC mortality in the exposed and unexposed groups compared to the control group, respectively. Given ITT analysis, the final part of the equation (S-2) gives the
estimate of first adjusted RR \((aRR_1)\) taken as the average of \(RR_E\) and \(RR_E\)
weighted by the screening rate.

\[
aRR_1 \equiv RR_E \times r_E + RR_E \times (1 - r_E) \tag{S-3}
\]

Note that \(P(D|\bar{I})\) represents the expected mortality rate of CRC in the absence
of screening, equivalent to the mortality rate of CRC of the uninvited (control) group
in the randomized controlled trial. In population-based FIT service screening, it is
impossible to have the uninvited group as seen in the RCT, the pre-screening group
with the adjustment for annual natural growth rate of increasing incidence rate of
CRC as mentioned in the text of the statistical part in methods section was used for a
proxy for the risk of being dead from CRC in the invited group.

**Relative rate of advanced CRC or CRC death with the full adjustment from uptake
screening until the completeness of colonoscopy**

Figure 1 in the main text shows a cascade of processes from uptake screening,
the referral for confirmatory diagnostic examination until complete colonoscopy in
population-based FIT service screening given the population is invited. The invited
group in the equation (S-1) was further classified into six groups.

Therefore, the numerator of equation (S-1) on the risk for being dead from CRC
in the invited group is first divided into two parts as in the equation (S-2), the exposed
group and the unexposed group. The exposed group part is further decomposed into five corresponding parts. The expression is written as:

\[
P(D|I) = P(D, E, Po, Re, Cp, Ci|I) + P(D, E, Po, Re, Cp, Ci\bar{|}|I) + P(D, E, Po, Re, \bar{Cp}|I) + P(D, E, \bar{Po}|I) + P(D, E\bar{|}|I)
\]

(S-4)

The first five parts on the right side of equation are the decomposition of the conditional probability \(P(D, E|I)\) for those who exposed to screening (E) after invitation (I) in opposite to the counterpart \(P(D, E\bar{|}|I)\) for the unexposed group \(E\bar{|}\) after invitation (I). Note that as whether to have complete colonoscopy to reach the cecum can only be assessed conditioned on the fact that positive-test subjects were referred to colonoscopy after they had the uptake of FIT screening, all these conditional probabilities are further expanded in a forward manner following a cascade of subsequent dichotomous outcomes for positive FIT test (Po), referral (Re), colonoscopy (Cp), and complete colonoscopy (Ci). This means that the conditional probability is further expanded given the positive result of FIT. Once the negative outcome (denoted by \(Po\)) is found, the expansion of the conditional probability ends.

Furthermore, as all these conditional probabilities refer to a cascade of the abovementioned characteristics and its associated death from CRC given those who were invited to screen they can be simplified by memoryless property for the outcome.
earlier. Take \( P(D, E, P_0, R_e, C_p, C_i | I) \) as an example, it can be re-expressed by the following conditional probabilities.

\[
P(D, E, P_0, R_e, C_p, C_i | I) \\
= P(D | E, P_0, R_e, C_p, C_i, I) \times P(C_i | E, P_0, R_e, C_p, I) \times P(C_p | E, P_0, R_e, I) \times \\
P(R_e | E, P_0, I) \times P(P_0 | E, I) \times P(E | I)
\]

Death from CRC given complete colonoscopy or not is therefore independent of events earlier. This means once information on whether to reach cecum with colonoscopy is exactly known the previous outcomes on positive FIT test, the referral to have confirmatory diagnosis, the administration of colonoscopy, cannot provide additional information. Namely, \( P(D | E, P_0, R_e, C_p, C_i, I) = P(D | C_i) \). Recall that such a property is called conditional independence. The same idea is also applied to other parts. Equation (S-5) can then be rewritten as

\[
= P(D | C_i) \times P(C_i | C_p) \times P(C_p | R_e) \times P(R_e | P_0) \times P(P_0 | E) \times P(E | I)
\]

Let \( r_{CI} \), \( r_{CP} \), \( r_{REF} \), \( r_{POS} \), and \( r_E \) denote cecal intubation rate (CIR), the proportion of selecting colonoscopy, referral rate to diagnostic examination, positive rate, and screening rate. Equation (S-6) can be further expressed as

\[
= P(D | C_i) \cdot r_{GI} \cdot r_{CP} \cdot r_{REF} \cdot r_{POS} \cdot r_E
\]
Equation (S-4) can be expressed as

\[
P(D|I) = P(D|CI) \cdot r_{CI} \cdot r_{CP} \cdot r_{REF} \cdot r_{POS} \cdot r_E + \\
P(D|CI) \cdot (1-r_{CI}) \cdot r_{CP} \cdot r_{REF} \cdot r_{POS} \cdot r_E + \\
P(D|CP) \cdot (1-r_{CP}) \cdot r_{REF} \cdot r_{POS} \cdot r_E + \\
P(D|RE) \cdot (1-r_{REF}) \cdot r_{POS} \cdot r_E + \\
P(D|PO) \cdot (1-r_{POS}) \cdot r_E + P(D|E) \cdot (1-r_E)
\]

(S-7)

To render the equation of (S-7) adapted for the regression model and amenable to estimation of parameters, taking into account age, gender, and increasing incidence trend, we applied the Bayesian DAG Poisson regression model proposed by Wu et al\(^1\) for evaluating the effectiveness of Taiwanese Nationwide CRC Screening Program. The Bayesian DAG Poisson regression model links the relationship of the status of exposure to screen (1=exposed for attenders, 0=unexposed for non-attenders) with the number of advanced-stage CRC or CRC death in each group that is assumed to follow Poisson distribution. The self-selection bias using adjusted RR (aRR\(_1\)) in light of ITT analysis can be adjusted by the following Poisson regression model expressed by

\[
\log(\mu) = \log(PY) + \beta_B + \sum_{i=1}^{8} \beta_i X_i
\]

(S-8)

where \(\mu\) denotes the expected number of advanced-stage CRC or CRC death, \(PY\) is
the corresponding person years, X₁ − X₆ are indicator variables for the six groups in
the invited group, X₇ for age, and X₈ for sex. Note that β₇ accounts for the natural
growth rate of incidence trend of CRC during the screening period had screening not
taken place and was estimated as 0.0443 (se=0.000243) based on the extrapolation
with time trend before the screening period.

Taking the exponent of six regression coefficients (β₁ − β₆), the corresponding
regression coefficients, gives six RRs for the detailed groups compared to the
uninvited group.

The self-selection bias made by using adjusted RR (aRR₂) in light of ITT
analysis, making allowance for positive FIT test, referral, the choice of colonoscopy,
and complete colonoscopy, was formulated as calculated as

\[
aRR₂ = e^{β₁} \cdot r_{CI} \cdot r_{CP} \cdot r_{REF} \cdot r_{POS} \cdot r_E + e^{β₂} \cdot (1 - r_{CI}) \cdot r_{CP} \cdot r_{REF} \cdot r_{POS} \cdot r_E
\]

\[
+ e^{β₃} \cdot (1 - r_{CP}) \cdot r_{REF} \cdot r_{POS} \cdot r_E + e^{β₄} \cdot (1 - r_{REF}) \cdot r_{POS} \cdot r_E
\]

\[
+ e^{β₅} \cdot (1 - r_{POS}) \cdot r_E + e^{β₆} \cdot (1 - r_E)
\]

(S-9)

**Sensitivity Analysis**

In addition to estimating the adjusted relative rate as above, the equation (S-9)
enables us to further calculate the relative rate of advanced-stage CRC or CRC death
for the conditions:
• if all subjects with positive FIT results complying to diagnostic examination 
(100% referral rate, $r_{REF} = 1$) (no group 4),

$$aRR_{REF} = e^{\beta_1 \cdot r_{CI} \cdot r_{CP} \cdot r_{POS} \cdot r_E} + e^{\beta_2 \cdot (1 - r_{CI}) \cdot r_{CP} \cdot r_{POS} \cdot r_E} + e^{\beta_3 \cdot (1 - r_{CP}) \cdot r_{POS} \cdot r_E} + e^{\beta_2 \cdot (1 - r_{POS}) \cdot r_E} + e^{\beta_6 \cdot (1 - r_E)}$$

• if all subjects with positive FIT results complying to diagnostic examination with colonoscopy (100% choice of colonoscopy, ($r_{REF} = 1$ and $r_{CP} = 1$)) (without groups 3 and 4),

$$aRR_{REF} = e^{\beta_1 \cdot r_{CI} \cdot r_{POS} \cdot r_E} + e^{\beta_2 \cdot (1 - r_{CI}) \cdot r_{POS} \cdot r_E} + e^{\beta_5 \cdot (1 - r_{POS}) \cdot r_E} + e^{\beta_6 \cdot (1 - r_A)}$$

• and if all subjects with positive FIT results complying to diagnostic examination with complete colonoscopy till cecum (100% complete colonoscopy, ($r_{REF} = 1$, $r_{CP} = 1$ and $r_{CI} = 1$) (without groups 2, 3 and 4),

$$aRR_{REF} = e^{\beta_1 \cdot r_{POS} \cdot r_E} + e^{\beta_5 \cdot (1 - r_{POS}) \cdot r_E} + e^{\beta_6 \cdot (1 - r_E)}.$$

Bayesian Monte Carlo Markov Chain (MCMC) method was implemented to estimate the adjusted RRs as indicated above for the comparisons of outcome in both the exposed and the unexposed group with the comparator using the pre-screening epoch between 1998 and 2003, the period before nationwide CRC screening was launched, taking into account age, gender and the growth rate of incidence trend in the
absence of screening. Supplementary Tables 1 and 2 show the data layout used for the Poisson regression model.

In the framework of Bayesian DAG Poisson regression model, we assigned the screening rate \( r_E \) following a Beta distribution, \( \text{Beta}(3067853, 2349846) \), where the former and latter numbers represent numbers of the exposed and the unexposed group in this study. \( \beta_0 \) follows normal distribution, \( \text{N}(0.0443, 0.000243) \). We used the non-informative priors for \( \beta_1 - \beta_4 \), which follow normal distribution, \( \text{N}(0, 10^{-6}) \).

For the adjustment of stage shifting, because cancer stage information was insufficient before 2003, we made use of the stage information from a study conducted by Ju et al. at Taipei Veterans General Hospital and dataset from National Taiwan University Hospital during the period of 1991 and 2000 to derive the proportion of advanced-stage CRC (AJCC stage II and higher) as 86% with the Bayesian conjugate approach.\(^2\) (Supplementary Table 3) Accordingly, the estimated number of advanced-stage CRC was 2,387. The information on anatomical site of CRC was derived in the same way and the proportion of distal cancer was 78%. (Supplementary Table 4) Information on CRC death from distal cancer was also derived in the same manner and specified in Supplementary Table 5.
### Supplementary Table 1

Tabular data for CRC death by the exposure to screen and the control group cross-tabulated by sex and age groups

<table>
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<th>Age Group</th>
<th>Number of death</th>
<th>Person years</th>
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<td></td>
<td>Exposed</td>
<td>Unexposed</td>
</tr>
<tr>
<td>Male</td>
<td></td>
<td></td>
</tr>
<tr>
<td>50-54</td>
<td>45</td>
<td>680</td>
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<tr>
<td>55-59</td>
<td>259</td>
<td>1,710</td>
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<td>60-64</td>
<td>401</td>
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<tr>
<td>65-69</td>
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<tr>
<td>70+</td>
<td>514</td>
<td>3,192</td>
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<tr>
<td>Female</td>
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<td></td>
</tr>
<tr>
<td>50-54</td>
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<td>513</td>
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<tr>
<td>55-59</td>
<td>299</td>
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<td>337</td>
<td>1,144</td>
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<tr>
<td>70+</td>
<td>374</td>
<td>2,204</td>
</tr>
<tr>
<td>Total</td>
<td>3,077</td>
<td>15,550</td>
</tr>
</tbody>
</table>
Supplementary Table 2

Tabular data for CRC advanced CRC by the exposure to screen and the control group cross-tabulated by sex and age groups

<table>
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<th></th>
<th>Number of advanced CRC</th>
<th>Person years</th>
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<td>Exposed</td>
<td>Unexposed</td>
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<td></td>
<td></td>
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<tr>
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<tr>
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<td>2,572</td>
</tr>
<tr>
<td><strong>Total</strong></td>
<td>6,381</td>
<td>23,689</td>
</tr>
</tbody>
</table>
**Supplementary Table 3**

Number of advanced CRC and the proportion of cancers with stage II or higher during the period of 1991 to 2000

<table>
<thead>
<tr>
<th>Cohort / Dataset</th>
<th>N</th>
<th>Number of stage II+ colorectal cancer cases</th>
<th>%</th>
</tr>
</thead>
<tbody>
<tr>
<td>NTUH</td>
<td>169</td>
<td>165</td>
<td>98</td>
</tr>
<tr>
<td>TVGH</td>
<td>3230</td>
<td>2746</td>
<td>85</td>
</tr>
</tbody>
</table>

NTUH: National Taiwan University Hospital  
TVGH: Taipei Veteran General Hospital  
The proportion of stage II and higher colorectal cancer was approximately 86% based on meta-analysis and Bayesian approach

**Supplementary Table 4**

Number of distal CRC and its proportion among all incident CRC in the period of 1991 to 2000

<table>
<thead>
<tr>
<th>Cohort / Dataset</th>
<th>N</th>
<th>Number of distal colorectal cancers</th>
<th>%</th>
</tr>
</thead>
<tbody>
<tr>
<td>NTUH</td>
<td>166</td>
<td>114</td>
<td>67</td>
</tr>
<tr>
<td>TVGH</td>
<td>3230</td>
<td>2552</td>
<td>79</td>
</tr>
</tbody>
</table>

NTUH: national Taiwan University Hospital  
TVGH: Taipei Veteran General Hospital  
The proportion of distal CRC was approximately 78% based on meta-analysis and Bayesian approach

**Supplementary Table 5**

Number of distal CRC death and its proportion among all CRC deaths in the period of 1981 to 2000

<table>
<thead>
<tr>
<th>Cohort / Dataset</th>
<th>Colorectal cancer death, n</th>
<th>Distal colorectal cancer death, n</th>
<th>%</th>
</tr>
</thead>
<tbody>
<tr>
<td>NTHU</td>
<td>105</td>
<td>70</td>
<td>67</td>
</tr>
<tr>
<td>TVGH</td>
<td>3143</td>
<td>2462</td>
<td>78</td>
</tr>
</tbody>
</table>

NTUH: national Taiwan University Hospital  
TVGH: Taipei Veteran General Hospital  
The proportion of death from distal site of CRC among all CRC death was approximately 77% based on meta-analysis and Bayesian approach
References
